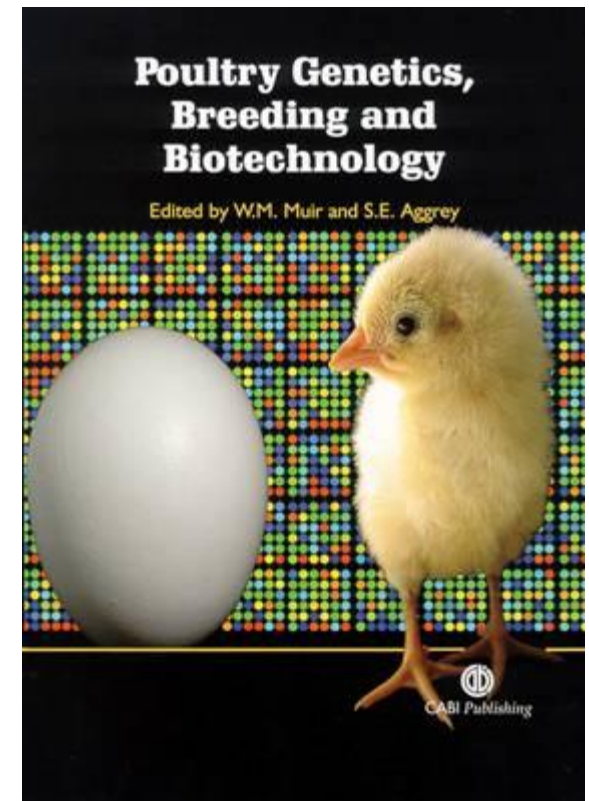


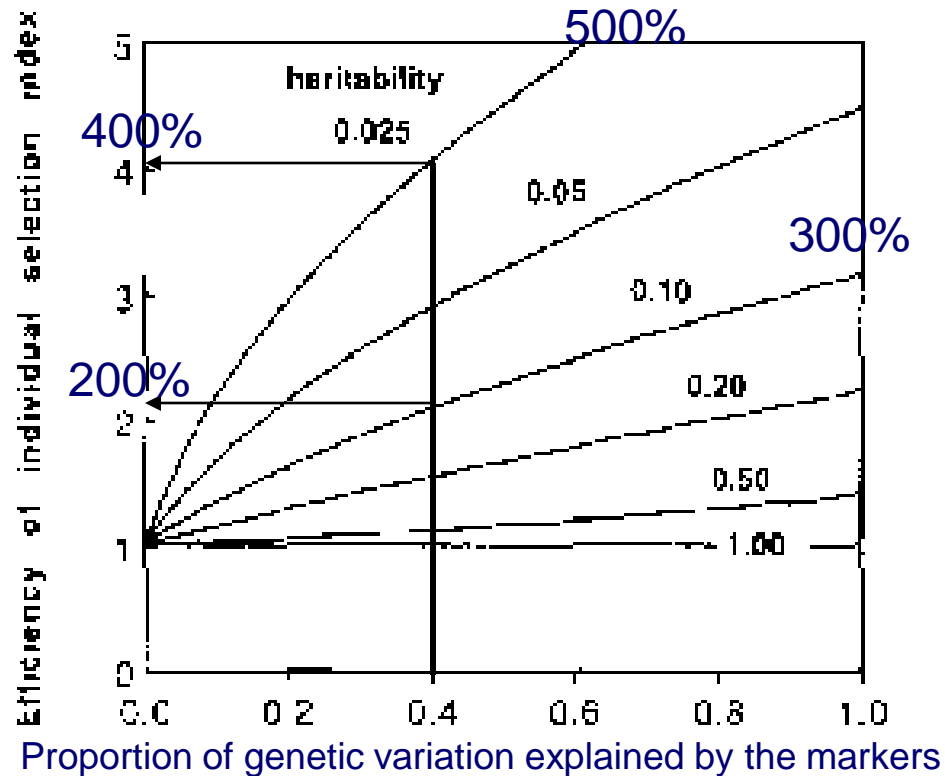


# Genomic Selection: Promise And Concerns

William Muir  
Purdue University



# Efficiency of MAS Relative to Individual Selection



Lande and Thompson (1990)

Problem: Cannot identify QTL that account for even moderate proportion of genetic variation with lowly heritable traits

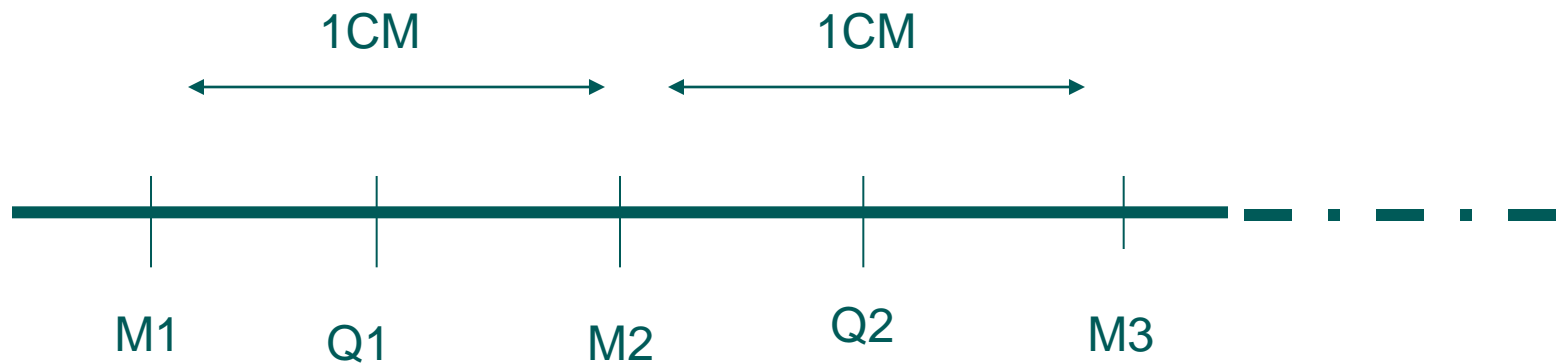
# Detection of QTL

- Also requires phenotypic data
- Low Heritability Traits
  - Cannot Establish Marker-QTL associations
  - MAS is Most Needed for Such Traits

PARADOX

# Potential Solution

- Genome Wide MAS (GMAS)
  - Meuwissen et al 2001
  - Uses all Markers
  - Dense
    - Every 1cM



# Implementation (Meuwissen et al, 2001)

- Combine All Data In Mixed Model
  - QTL effect Assumed Sampled From Distribution
  - Random Effect

Estimation

$$Y = XB + ZG + \varepsilon$$

where  $G_i$  represents the genetic effect of the  $i^{\text{th}}$  haplotype,  $Z_i$  is an incidence matrix and has a 0, 1, 2 for the number of haplotypes of type  $G_i$  present in the  $j^{\text{th}}$  animal

Prediction

$$\hat{Y}_j = GEBV = \sum_i Z_i \hat{G}_i$$

# Estimation

- Requires Multiple Generations of Data
  - All Individuals
    - Genotyped at All Loci (Z matrix)
    - Phenotypes Measured for All Traits (Y Matrix)
      - As Many Traits as Desired
    - Fixed Effects (Age, Sex, Block, etc)
      - Recorded
      - (X Matrix)

$$Y = XB + ZG + \varepsilon$$

# Maximum Likelihood Estimation of B and G

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}\mathbf{X}' & \mathbf{Z}'\mathbf{Z} + \frac{\sigma_e^2}{\sigma_g^2} \mathbf{I} \end{bmatrix} \begin{bmatrix} \boldsymbol{\beta} \\ \mathbf{G} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{Y} \\ \mathbf{Z}'\mathbf{Y} \end{bmatrix}$$

Assumes that genetic variance associated with each marker is equal to the total genetic variance divided by the total number of markers (Bayesian Ridge Regression)

# How Well Does it Work?

Compare Accuracy of  
Genomic EBV (GEBV) With  
BLUP EBV (BEBV)  
Using Gene Level Simulations



# Assumptions are Critical

- All MAS requires Linkage Disequilibrium (LD)
  - What is LD
  - How is it generated

# Linkage Disequilibrium

Genotype/Haplotype		Frequency
$\begin{array}{c c} m^A & Q^B \\ \hline m^a & Q^b \end{array}$	Coupling Phase	$p(AB/ab)$
$\begin{array}{c c} m^A & Q^b \\ \hline m^a & Q^B \end{array}$	Repulsion Phase	$p(Ab/aB)$

$$LD = p(AB/ab) - p(Ab/aB)$$

Any factor changing the relative frequency of coupling vs. repulsion phase impacts LD

# LD Generation

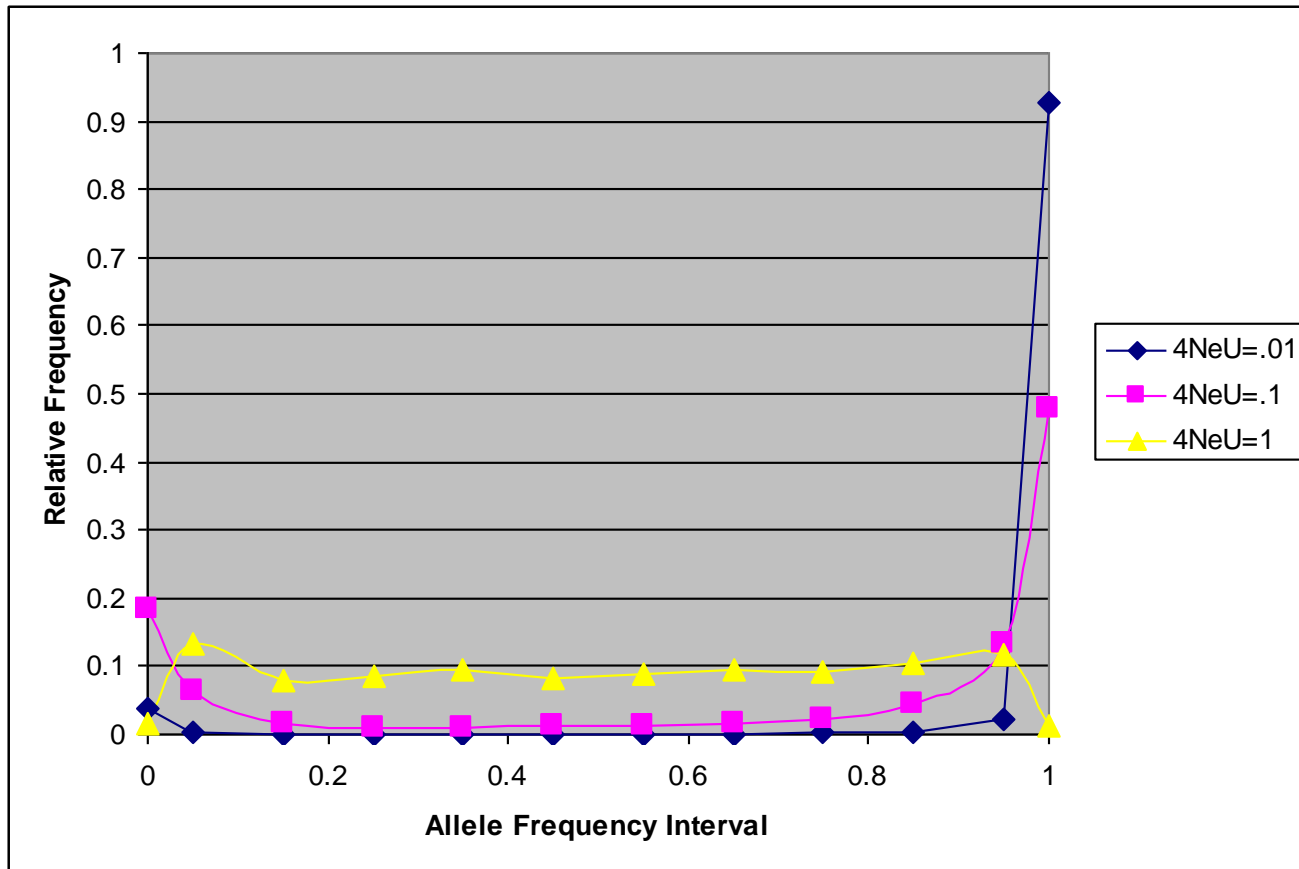
- All forces that change allele frequencies
  - Mutation
  - Migration
  - Selection
  - Genetic Drift

# Starting Conditions

- Hardy-Weinberg Equilibrium
  - Within and between loci
  - Generate LD by Random Drift
- Mutation-Drift Equilibrium
  - Pre-existing LD
  - LD has not decayed from original mutation event
  - Dependent on mutation rates (types of markers) and population size

# Mutation Drift Equilibrium (MDE)

$N_e=100$  Generations=1000



$N_e=100$

QTL

$u=2.5 \times 10^{-5}$

$u=2.5 \times 10^{-4}$

$u=2.5 \times 10^{-3}$

Markers

Markers=Uniform  
Distribution  
all Polymorphic

QTL=4%

Polymorphic  
Steady State

# Generations of Training Duration and Accuracy of Prediction

	Generation	Genotype	Phenotype	Predict Breeding Value	Accuracy
Training Generations	0	Z	and	$\hat{Y}_0$	$r_0$
	1	Z		$\hat{Y}_1$	$r_1$
	...	Z		$\hat{Y}_t$	$r_t$
	t	Z		$\hat{Y}_t$	$r_t$
Prediction Generations	t+1	Z	No Phenotypic Data	$\hat{Y}_{t+1}$	$r_{t+1}$
	t+2	Z		$\hat{Y}_{t+2}$	$r_{t+2}$
	...	Z		$\hat{Y}_{t+n}$	$r_{t+n}$
	t+n	Z		$\hat{Y}_{t+n}$	$r_{t+n}$

Accuracy=Correlation Between Predicted and True Breeding Value

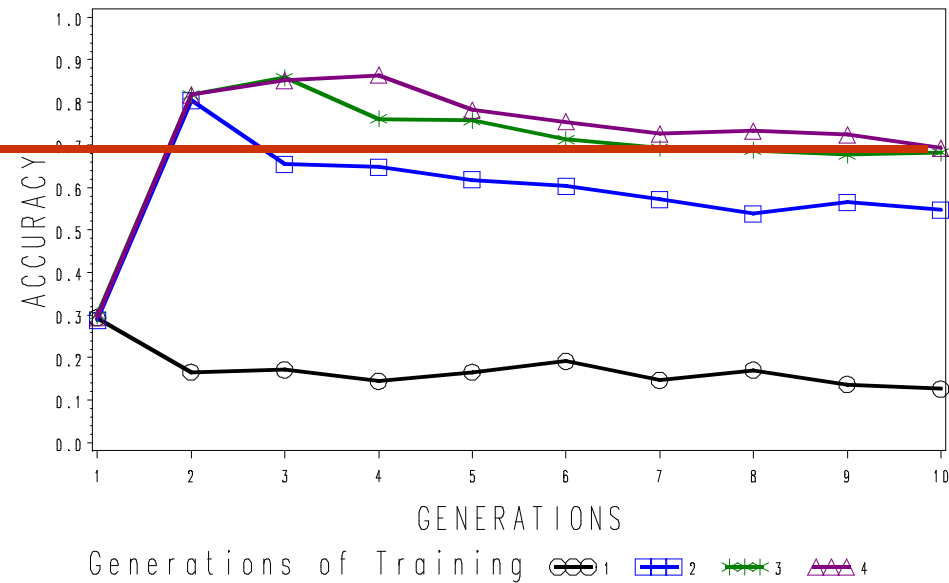
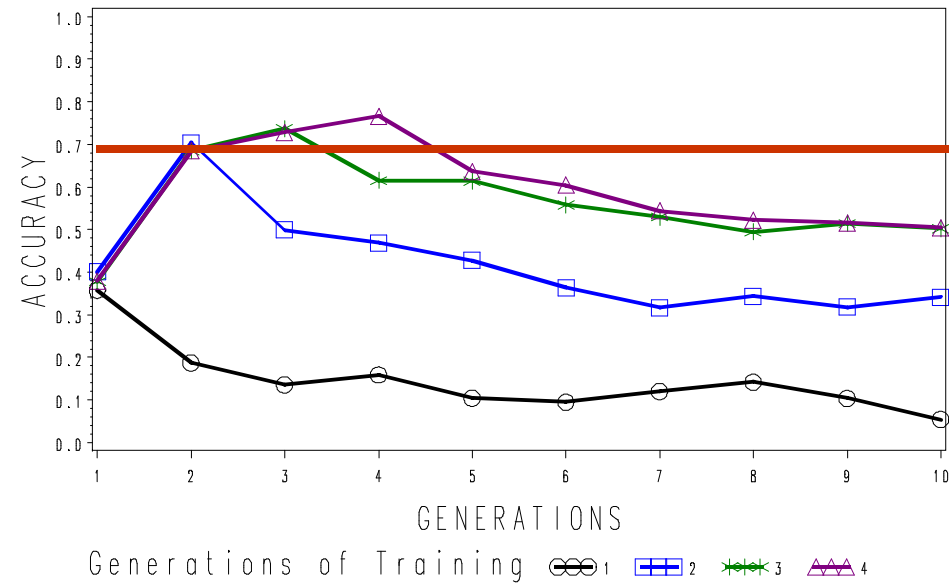
# Results

# GEBV

## Accuracy with $h^2=.5$

Starting in HWE

Starting in MDE



$N=128, N_e=16$

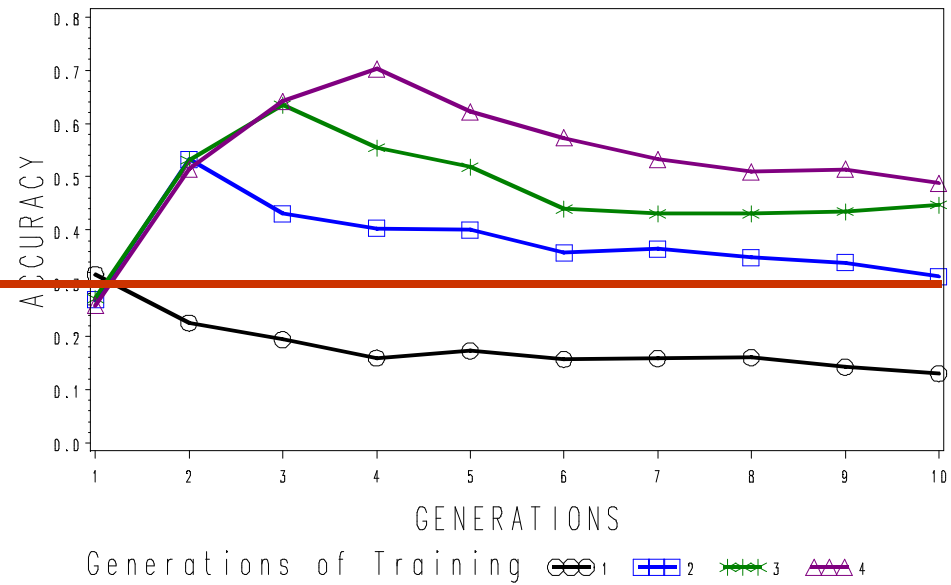
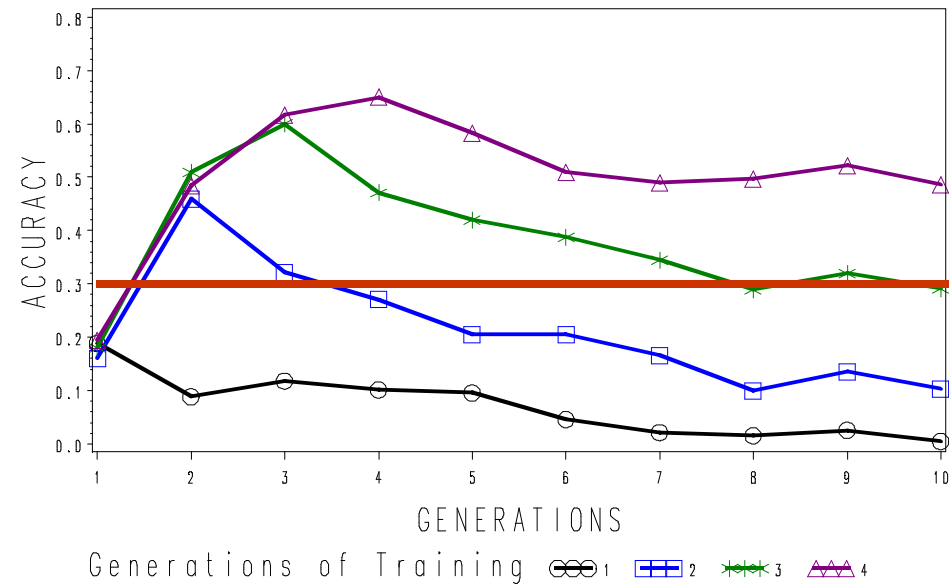


# GEBV

## Accuracy with $h^2=.1$

Starting in HWE

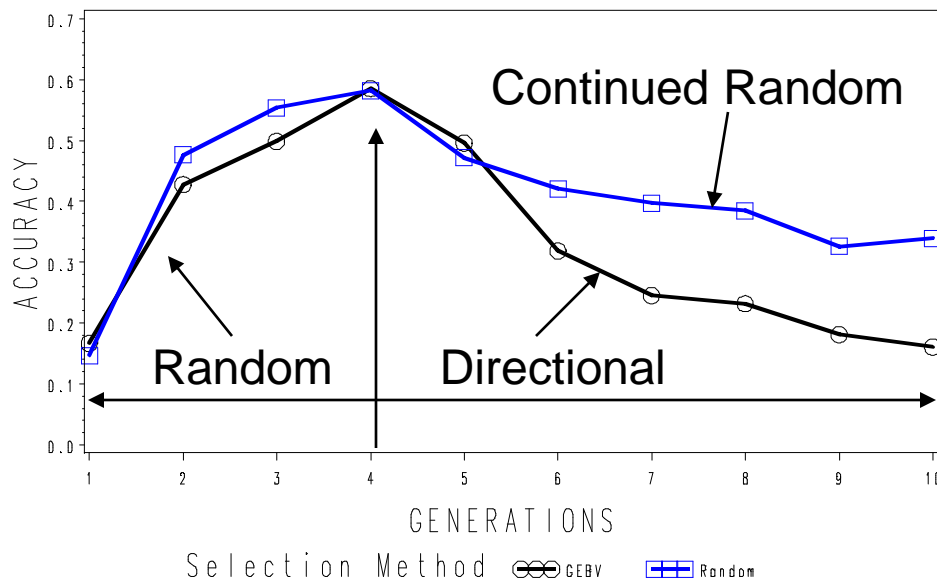
Starting in MDE



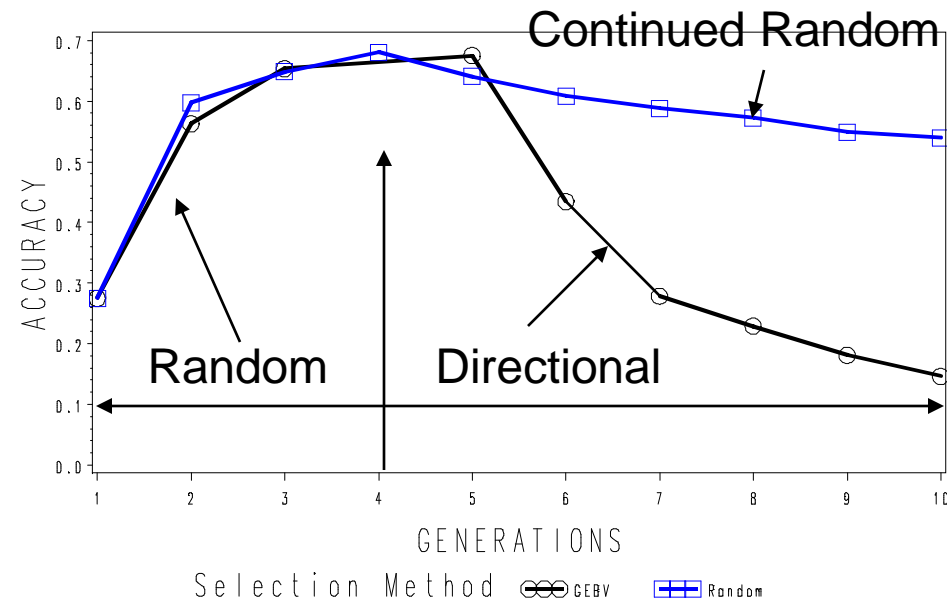
$N=128, N_e=16$

# Effect of Random vs. Directional Selection on Accuracy

Starting in HWE



Starting in MDE



$h^2=.1$   $N=256$ ,  $N_e=32$ , 100/100 Marker/QTL loci distributed on 100cM.  
 (average over 60 replicates, SEM=.02).

# Alternative Approach

## Genomic Relationship Matrix

- Assumes
  - Dense markers evenly spaced across the genome
  - Assumes markers are in LD with QTL affecting trait(s) of interest
  - Alike in State (AIS) alleles were at one time a result of a single mutation, thus IBD when traced back in evolutionary time
  - Each marker account for an equal proportion of genetic variance (infinitesimal model)
  - Genetic Effects are Normally Distributed

# Compute (AIS) relationship matrix (G)

$$TA_k = 2 \frac{\sum_{i=1}^2 \sum_{j=1}^2 I_{ij}}{4}$$

TA<sub>k</sub>=total allelic relationship at kth locus  
 TA<sub>k</sub>=2x coefficient of relationship  
 (Malecot. 1948)

$$G_{xy}^* = \frac{\sum_{k=1}^L TA_k}{L} = \frac{2 \sum_{i=1}^L \left( \frac{\sum_{j=1}^2 \sum_{j=1}^2 I_{ij}}{4} \right)}{L}$$

$$\mathbf{G} = \sigma_{A^*}^2 \mathbf{G}^*$$

$$\sigma_{A^*}^2$$

Is the additive genetic variance associated with the markers for the trait

$$\sigma_{A^*}^2 < \sigma_A^2$$

Note: with low marker density the markers may not capture any genetic variance

Individual	LOCUS										Pedigree
	A		B		C		D		E		
1	2	2	1	1	1	2	1	1	2	2	
2	1	2	1	2	2	2	1	2	1	1	
3	1	2	1	1	1	2	1	2	1	2	
4	2	2	1	1	2	2	1	1	2	1	
5	2	1	1	2	2	2	1	1	2	1	
6	2	2	1	1	2	2	1	1	2	1	

Individuals (X,Y)						Total	rx <sub>y</sub>
x=1						9	1.8
y=1							
sum		4	4	2	4		
shared alleles		2	2	1	2		
x=1						4	0.8
y=2							
sum		2	2	2	2		
shared alleles		1	1	1	1		

	AIS		G=GRM				IBD		PEDIGREE	A	
	1	2	3	4	5	6					
1	1.8	0.8	1.2	1.6	1.2	1.6	0.5	0.5	0.5	0.5	0.5
2	0.8	1.4	1	1.2	1.2	1.2	0.5	0.5	0.5	0.5	0.5
3	1.2	1	1.2	1.2	1.8	1.2	0.5	0.5	1	0.5	0.5
4	1.6	1.2	1.2	1.8	1.4	1.8	0.5	0.5	0.5	1	0.5
5	1.2	1.2	1	1.4	1.4	1.4	0.5	0.5	0.5	0.5	1
6	1.6	1.2	1.2	1.8	1.4	1.8	0.5	0.5	0.5	0.5	0.5

Parents assumed not related (False)

Parents assumed non inbred (false)

Full sibs assumed = relationship (false)

# G\* Computed Directly from M

Individual	LOCUS										22=1 12=0 11=-1		
	A		B		C		D		E				
	1	2	1	2	1	2	1	2	1	2			
1	2	2	1	1	1	2	1	1	2	2			
2	1	2	1	2	2	2	1	2	1	1			
3	1	2	1	1	1	2	1	2	1	2			
4	2	2	1	1	2	2	1	1	2	1			
5	2	1	1	2	2	2	1	1	2	1			
6	2	2	1	1	2	2	1	1	2	1			
			<b>M</b>	N individuals x p markers					<b>M'</b>	p markers x N individuals			
1	1	-1	0	-1	1		1	0	0	1	0	1	
2	0	0	1	0	-1		-1	0	-1	-1	0	-1	
3	0	-1	0	0	0		0	1	0	1	1	1	
4	1	-1	1	-1	0		-1	0	0	-1	-1	-1	
5	0	0	1	-1	0		1	-1	0	0	0	0	
6	1	-1	1	-1	0								
	0.8	-0.2	0.2	0.6	0.2	0.6		1.8	0.8	1.2	1.6	1.2	1.6
	-0.2	0.4	0	0.2	0.2	0.2		0.8	1.4	1	1.2	1.2	1.2
	0.2	0	0.2	0.2	0	0.2		1.2	1	1.2	1.2	1	1.2
	0.6	0.2	0.2	0.8	0.4	0.8		1.6	1.2	1.2	1.8	1.4	1.8
	0.2	0.2	0	0.4	0.4	0.4		1.2	1.2	1	1.4	1.4	1.4
	0.6	0.2	0.2	0.8	0.4	0.8		1.6	1.2	1.2	1.8	1.4	1.8
			<b>MM'/5</b>							<b>G*</b>			
			dimension nxn							Note that $1+MM'/5=G^*$			
										The are the same			

# Mixed Model Equations

$$\begin{bmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{b} \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{Y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Y} \end{bmatrix}$$

Simplifications If  $\mathbf{R} = \mathbf{I}\sigma_e^2$

$$\mathbf{G} = \sigma_{A^*}^2 \mathbf{M}'\mathbf{M} / L$$

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}'\mathbf{X} & \mathbf{Z}'\mathbf{Z} + \frac{\sigma_e^2}{\sigma_{A^*}^2} (\mathbf{M}'\mathbf{M} / L)^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{b} \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{Y} \\ \mathbf{Z}'\mathbf{Y} \end{bmatrix}$$

Note one can multiply both sides of the second equation by  $\mathbf{G}$  to avoid an inverse

## Example

$$\mathbf{Y} = \begin{bmatrix} 7 \\ 9 \\ 10 \\ 6 \\ 9 \\ 11 \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \quad b = [\mu_0] \quad \mathbf{Z} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix} \quad \mathbf{u} = \begin{bmatrix} u_1 \\ u_2 \\ u_3 \\ u_4 \\ u_5 \\ u_6 \end{bmatrix}$$

$$\mathbf{M} = \begin{bmatrix} 1 & -1 & 0 & -1 & 1 \\ 0 & 0 & 1 & 0 & -1 \\ 0 & -1 & 0 & 0 & 0 \\ 1 & -1 & 1 & -1 & 0 \\ 0 & 0 & 1 & -1 & 0 \\ 1 & -1 & 1 & -1 & 0 \end{bmatrix} \quad \mathbf{MM}' / L = \begin{bmatrix} .8 & -.2 & .2 & .6 & .2 & .6 \\ -.2 & .4 & 0 & .2 & .2 & .2 \\ .2 & 0 & .2 & .2 & 0 & .2 \\ .6 & .2 & .2 & .8 & .4 & .8 \\ .2 & .2 & 0 & .4 & .4 & .4 \\ .6 & .2 & .2 & .8 & .4 & .8 \end{bmatrix} \quad \begin{aligned} \sigma_A^2 &= 10 \\ \sigma_{A^*}^2 &= 5 \\ \sigma_\varepsilon^2 &= 20 \end{aligned}$$

Note, only ½ the additive genetic variance was captured by the markers 24



Y				Z																
7		1	0	0	0	0	0	0	V(A)=	5										
9		0	1	0	0	0	0	0	V(E)=	20										
10		0	0	1	0	0	0	0												
6		0	0	0	1	0	0	0		XX										
9		0	0	0	0	1	0	0		6										
11		0	0	0	0	0	1	0												
X				MM/5						X'Y										
1		0.8	-0.2	0.2	0.6	0.2	0.6	0.6		52										
1		-0.2	0.4	0	0.2	0.2	0.2	0.2												
1		0.2	0	0.2	0.2	0	0.2	0.2												
1		0.6	0.2	0.2	0.8	0.4	0.8	0.8												
1		0.2	0.2	0	0.4	0.4	0.4	0.4												
1		0.6	0.2	0.2	0.8	0.4	0.8	0.8												
V(A*)G*ZX				V(A*)GZZ+V(E)I						V(A*)GZY										
11		24	-1	1	3	1	3	3		89										
4		-1	22	0	1	1	1	1		37										
4		1	0	21	1	0	1	1		34										
15		3	1	1	24	2	4	4		126										
8		1	1	0	2	22	2	2		68										
15		3	1	1	4	2	24	24		126										
6	1	1	1	1	1	1	1	b		52										
11	24	-1	1	3	1	3	3	u1		89										
4	-1	22	0	1	1	1	1	u2		37										
4	1	0	21	1	0	1	1	u3		34										
15	3	1	1	24	2	4	4	u4		126										
8	1	1	0	2	22	2	2	u5		68										
15	3	1	1	4	2	24	24	u6		126										
				LHS						RHS										
b		0.238268	-0.00798	-0.01022	-0.01037	-0.00629	-0.00886	-0.00629		52										8.762384
u1		-0.07865	0.045659	0.005799	0.001807	-0.00248	0.001686	-0.00248		89										-0.25929
u2		-0.03389	0.003561	0.047264	0.001527	-0.00087	-0.00061	-0.00087		37										0.094488
u3		-0.03092	-0.00058	0.001379	0.04919	-0.00075	0.001505	-0.00075		34										-0.02194
u4	='	-0.11254	-0.00078	0.003063	0.003334	0.046655	0.001074	-0.00335		126										-0.1648
u5		-0.06107	0.000807	0.000747	0.003012	-0.0015	0.048432	-0.0015		68										-0.05796
u6		-0.11254	-0.00078	0.003063	0.003334	-0.00335	0.001074	0.046655		126										-0.1648

# Equivalent Model

## Estimation of Marker effects

$$\begin{bmatrix} \mathbf{X}'_{1,N} & \mathbf{X}_{N,1} & & \mathbf{X}_{1,N}' \mathbf{M}_{N,p} \\ \mathbf{M}'_{p,N} & \mathbf{X}_{N,1} & & \mathbf{M}'_{p,N} \mathbf{M}_{N,p} + \frac{\sigma_e^2}{\left(\frac{\sigma_{A^*}^2}{L}\right)} \mathbf{I} \end{bmatrix} \begin{bmatrix} \boldsymbol{\beta}_{1,1} \\ \mathbf{g}_{p,1} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'_{1,N} & \mathbf{Y}_{N,1} \\ \mathbf{M}'_{p,N} & \mathbf{Y}_{N,1} \end{bmatrix}$$

Dimension p x p

Each Marker effects is solved for

Model of Meuwissen et al (2001)

$$\hat{u}_i = GEBV_i = \mathbf{M}\mathbf{g} = \sum_j M_{ij} \hat{g}_j$$

							<b>M'</b>										<b>M</b>						
1	0	0	1	0	1		1	-1	0	-1	1		1	-1	0	-1	1		1	-1	0	-1	1
-1	0	-1	-1	0	-1																		
0	1	0	1	1	1																		
-1	0	0	-1	-1	-1																		
1	-1	0	0	0	0																		
							<b>MM'</b>																
3	-3	2	-3	1																			
-3	4	-2	3	-1																			
2	-2	4	-3	-1																			
-3	3	-3	4	-1																			
1	-1	-1	-1	2																			
							<b>I</b>																
1	0	0	0	0																			
0	1	0	0	0																			
0	0	1	0	0																			
0	0	0	1	0																			
0	0	0	0	1																			
							<b>LHS</b>																
6	3	-4	4	-4	0	<b>B</b>																	
3	23	-3	2	-3	1	g1																	
-4	-3	24	-2	3	-1	g2																	
4	2	-2	24	-3	-1	g3	='																
-4	-3	3	-3	24	-1	g4																	
0	1	-1	-1	-1	22	g5																	
							<b>inverse(LHS)</b>																
B		0.238268	-0.02055	0.030921	-0.03165	0.029414	0.002238	52															
g1		-0.02055	0.046795	0.002074	-0.00012	0.002068	-0.00194	24															
g2	='	0.030921	0.002074	0.047116	-0.00139	-0.00057	0.001958	-34	='														
g3		-0.03165	-0.00012	-0.00139	0.047031	0.00085	0.002119	35															
g4		0.029414	0.002068	-0.00057	0.00085	0.047091	0.002059	-33															
g5		0.002238	-0.00194	0.001958	0.002119	0.002059	0.045822	-2															
							<b>M</b>																
u1		1	-1	0	-1	1	-0.08491	-0.25929															
u2		0	0	1	0	-1	0.021935	0.094488															
u3	='	0	-1	0	0	0	0.012177	-0.02194	same as before														
u4		1	-1	1	-1	0	0.070134	-0.1648															
u5		0	0	1	-1	0	-0.08231	-0.05796															
u6		1	-1	1	-1	0		-0.1648															

**Genomic Selection in Poultry,  
Results with Broilers and  
Comparison with Traditional  
BLUP**

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- 
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- 
- **Collaborators:**
- Charles Pick and Tun-Ping Yu – DNA Landmarks, Montreal, Canada

# Potential Benefits

- Reduced Generation Interval
  - Select breeders as chicks
  - Housing and feed savings
- Increased accuracy
  - Low heritability traits
- Alternative for difficult or expensive traits to measure
  - Feed efficiency
  - Carcass composition
  - Disease resistance
  - Can address animal wellbeing concerns

# Design of experiment

- Lines
  - Male line
  - Female line
- Traits of selection
  - Breast Meat (high  $h^2$ )
  - Weight (medium  $h^2$ )

# Methods for estimation of breeding values (EBVs and GEBVs)

- EBV's
  - BLUP (multi-trait)
- GEBVs
  - GBLUP (multi-trait)
  - Single Step (ssGBLUP)
    - Aguilar, I., I. Misztal, D. L. Johnson, A. Legarra, S. Tsuruta *et al.* 2010 Hot topic: A unified approach to utilize phenotypic, full pedigree, and genomic information for genetic evaluation of Holstein final score. *Journal Of Dairy Science* **93**: 743-752.
    - Legarra, A., I. Aguilar, and I. Misztal, 2009 A relationship matrix including full pedigree and genomic information. *Journal Of Dairy Science* **92**: 4656-4663.
  - Advantages
    - Corrects for multi-trait selection bias
      - Vitezica, Z. G., I. Aguilar, I. Misztal, and A. Legarra, 2011 Bias in genomic predictions for populations under selection. *Genetics Research* **93**: 357-366.
    - Utilizes records from ungenotyped animals
      - Concurrent
      - Ancestral



# EBVs and GEBVs combined using an index

- Equal weight to all traits
- Standardized relative to additive variance (standard deviation)

# GBLUP and BLUP Training

- Two training generations
  - no selection
  - used historical phenotypes
  - Banked DNA samples
- Numbers genotyped (GBLUP)
  - 2,500 each line
- Phenotypes
  - BLUP and GBLUP
  - 280,000 each line

# Selection Program-GEBV

- Tiered
  - Tier 1
    - 800 **Genotyped** and Phenotyped
  - Tier 2
    - 200 Phenotyped
- Plays into strength of ssGBLUP
  - Uses all records
- Only birds from Tier 1 selected based on index
- Number selected
  - 20 Males
  - 200 Females

# Selection Program BLUP

- Tiered
  - Tier 1 (800) phenotyped
  - Tier 2 (200) phenotyped
- Only birds from Tier 1 selected based on index
- Number selected
  - 20 Males
  - 200 Females

# Duration

- 3 generations of selection
- Generation 4
  - Expanded to 4,000
  - Phenotyped only
    - Progeny test of Generation 3
    - Bases for comparison between methods

# Accuracy

- Correlation between the true and predicted breeding values
  - Don't know true EBV
- Equivalent formula
  - Legarra, A., C. Robert-Granie', E. Manfredi, and J. M. Elsen. 2008. Performance of genomic selection in mice. *Genetics* **180**: 611-618
  - Chen, C. Y., I. Misztal, I. Aguilar, A. Legarra, and W. M. Muir, 2011 Effect of different genomic relationship matrices on accuracy and scale. *Journal Of Animal Science* **89**: 2673-2679
  - Does not require true value of EBV to be known
  - Requires heritability to be known
  - Calculation of EBV does not include phenotypes of the generations they were estimated in

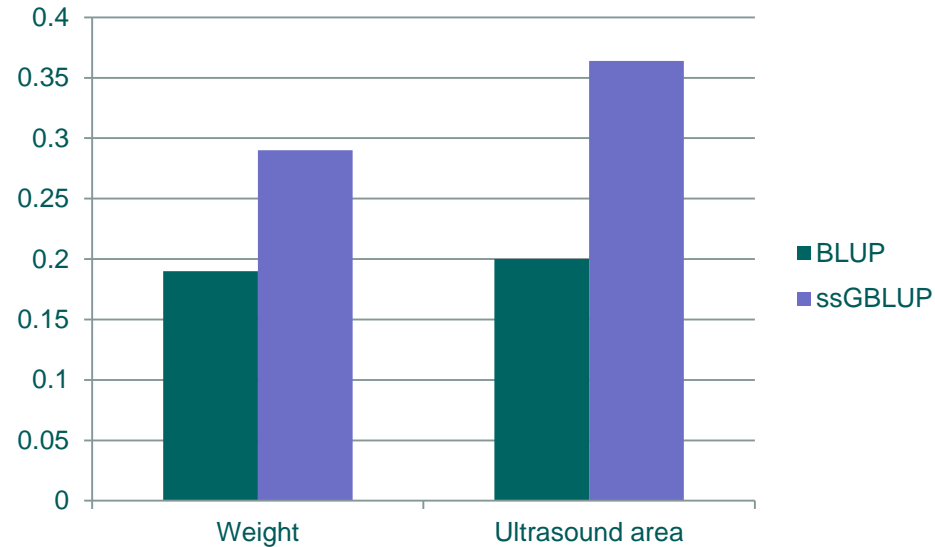
$$r(EBV, EB\hat{V})$$

$$\frac{r(EB\hat{V}, Y)}{h}$$

# Results

Generations 3-4

# Accuracy

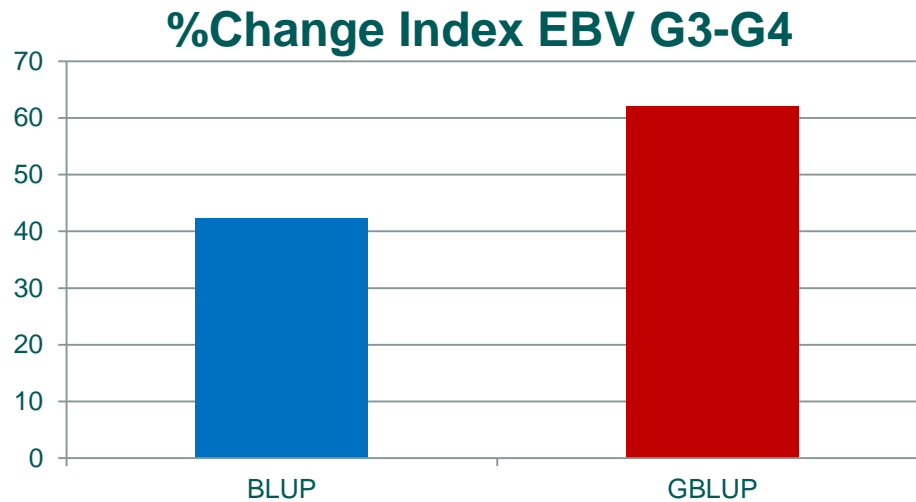


Chen, C. Y., I. Misztal, I. Aguilar, A. Legarra, and W. M. Muir, 2011 Effect of different genomic relationship matrices on accuracy and scale. *Journal Of Animal Science* **89**: 2673-2679



# Genetic Trend

## Female Line Index



# Genetic Trend Male Line Index

**%Change Index EBV G3-G4**



# Challenges

## GBLUP lines

- Some mismatch in early generations between DNA samples, pedigree, phenotypes

- Impact

$$\Delta G = \text{accuracy} * \text{intensity} * \sigma_a$$

- Reduced selection intensity in GBLUP lines

- Mismatches were random
    - Proportion of those selected was therefore at random

- Reduced accuracy in GBLUP program

- Training and SNP effects were updated each generation including current generation
    - Inaccurate SNP effects (phenotype- genotype relationships were incorrect)

# Opportunities

- Blood sampling on a large scale by farm workers in a chicken house.
  - Sample identification and tracking
  - Bar-coding, portable scanners and printers suitable for a chicken house
- Timely pedigree checking
  - Parent and offspring pedigree checking where possible
  - Full sib pedigree checking when parent information is not available
- Sample processing and genotyping to allow timely error checking
  - Timing is critical in chicken breeding (weeks not months)
  - Ability to keep selection candidates for pedigree verification if necessary
  - Space for storing selection candidates
  - Very different to cattle breeding..less space, less time

# Conclusions

- GBLUP more accurate than BLUP
- Genetic trends reflect increased accuracy of GBLUP
- GBLUP at a minimum was able to keep up with BLUP 😊
- Quality control essential for translation of technology to applications
- For some traits, especially those measured late in life, being as good as BLUP is good enough

# Conclusions

- Results are very encouraging
- Company will continue GBLUP program, particularly for
  - low heritability traits
  - Traits difficult or expensive to measure
  - Traits measured late in life cycle