



# Single-step GBLUP for populations under selection and with multibreed data

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ALIMENTATION  
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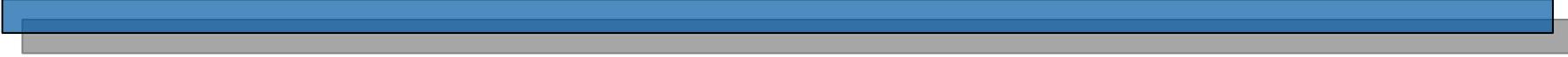
# ssGBLUP



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- ✓ A single-step approach was proposed based on **A** augmented with genomic information
- ✓ **Single step** has advantages such as simplicity, computational time and generality
- ✓ Most genotyped animals have undergone **strong selection**
  - → Single step needs adjustments (Aguilar et al., 2010, Chen et al., 2011, Forni et al. 2011)

# Effect of different genomic relationship matrices (**G**) on accuracy (Chen et al., 2011)<sup>1</sup>



280k broiler chicken

4k genotyped with 60k SNPs

Single-step GBLUP (**ssGBLUP**)

**G** was constructed using equal (0.5, **Geq**) or current (**GC**) allele frequencies

<sup>1</sup>Chen et al., 2011. *J Anim Sci*, 89: 2673-2679.

# Bias in genomic predictions

	Means of EBV			
	Genotyped		All	
	BW	BM	BW	BM
<b>BLUP</b>	2.77	2.27	1.67	1.38
<b>ssGBLUP</b>				
<b>GEq</b>	3.94	3.21	2.51	2.15
<b>GC</b>	0.45	0.32	-0.05	-0.09

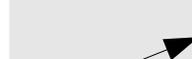
BLUP provides unbiased EBVs



EBVs are biased upwards



EBVs are biased downwards



**With bias, an accurate comparison of young and older animals is difficult !!**

# Bias in genomic predictions

**A22** and **G** matrices and their inverses were compared with regard to **AvgD** and **AvgOff** elements

	<b>AvgD</b>	<b>AvgOff</b>
<b>A22</b>	1.004	0.014
<b>G</b>		
<b>GC</b>	1.017	0

Unbiased evaluation in the ssGBLUP was obtained by adding a constant to **G** (=0.014) to be compatible with **A22**

# Data of litter size

(Forni et al., 2011)<sup>1</sup>



300k sows

2k genotyped with 60k SNPs

Single-step GBLUP (**ssGBLUP**)

Variance components estimation

<sup>1</sup> Fornin et al., 2011. *Gen Sel Evol*, 43: 1.

# Estimates of additive variance

Relationship matrix	Full data set	Genotyped subset
Pedigree ( <b>A</b> )	1.26 ±0.03	2.3 ±0.5
Genomic ( <b>G</b> )		
Equal gene freq	similar	3.5 ±0.6
Average min allele freq	similar	3.5 ±0.6
Realized gene freq	similar	2.4 ±0.4
Normalized ( <b>G</b> /trace( <b>G</b> ))	similar	2.3 ±0.3

Realistic estimates of additive variance



Parameter estimation may be biased, especially when data sets were small



# Statistics of different relationship matrices

Relationship matrix	Avg diagonal	Avg off-diagonal
Pedigree ( <b>A</b> )	1.0	0.03
Genomic ( <b>G</b> )		
Equal gene freq	1.25	0.59
Average min allele freq	1.70	1.02
Realized gene freq	0.94	0.00
Normalized ( $\mathbf{G}/\text{trace}(\mathbf{G})$ )	1.00	0.00

Parameter estimation may be biased if **G** is not compatible with **A**

# Why are the GEBVs biased ?

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- ✓ In traditional BLUP evaluation, all information about selection decisions is included in phenotypes and the relationship matrix ( $A$ ), and **no bias exists from selection** (Sorensen & Kennedy, 1984)
- ✓ Models for genomic prediction implicitly assume an **unselected genotyped population** (Hayes et al. 2009).
- ✓ In practice, genotyped individuals are highly selected, and genomic prediction models do not take into account this selection

How unbiased GEBVs can be predicted ?

# The single-step approach

It is based on the model  $\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}\mathbf{u} + e$

and  $p(\mathbf{u}) \sim N(0, \mathbf{H}\sigma_u^2)$

involves the genetic effect for nongenotyped ( $\mathbf{u}_1$ ) and genotyped ( $\mathbf{u}_2$ ) individuals. Here  $\mathbf{H}^{-1}$  is derived as

$$\mathbf{H}^{-1} = \begin{bmatrix} \mathbf{A}^{11} & \mathbf{A}^{12} \\ \mathbf{A}^{21} & \mathbf{A}^{22} + \mathbf{G}^{-1} - \mathbf{A}_{22}^{-1} \end{bmatrix}$$

where  $\mathbf{G}$  is a genomic relationship matrix



# Genomic relationship matrix $\mathbf{G}$

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Assume that  $\mathbf{G}$  is computed according to VanRaden 2008, using observed allelic frequencies

This implies that the average BV of genotyped individuals ( $\mathbf{u}_2$ ) is 0

$$p(\mathbf{u}_2) \sim N(0, \mathbf{G}\sigma_u^2)$$

This is possibly NOT the case if there is SELECTION

# An improved $\mathbf{G}$

Relative to the *pedigree base population*, the average BV of genotyped individuals ( $\mathbf{u}_2$ ), has a value possibly different from 0, say  $\mu$

$$p(\mu) \sim N(0, \alpha \sigma_u^2)$$

$\mu$  is random because of finite size (drift)

$$p(\mathbf{u}_2 | \mu) \sim N(\mathbf{1}\mu, \mathbf{G} \sigma_u^2)$$

$$p(\mathbf{u}_2 | \alpha) \sim N(0, (\mathbf{G} + \mathbf{1}\mathbf{1}'\alpha) \sigma_u^2)$$

$$p(\mathbf{u} | \alpha) \sim N(0, \mathbf{H}^\dagger \sigma_u^2)$$

where  $\mathbf{H}^\dagger$  is as  $\mathbf{H}$ , but substituting  $\mathbf{G}$  for  $\mathbf{G} + \mathbf{1}\mathbf{1}'\alpha$

# An improved $\mathbf{G}$

The mixed model equations are

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}'\mathbf{X} & \mathbf{Z}'\mathbf{Z} + \mathbf{H}^{\dagger-1}\lambda \end{bmatrix} \begin{bmatrix} \mathbf{b} \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}'\mathbf{y} \end{bmatrix}$$


$$\mathbf{H}^{\dagger-1} = \begin{bmatrix} \mathbf{A}^{11} & \mathbf{A}^{12} \\ \mathbf{A}^{21} & \mathbf{A}^{22} + (\mathbf{G} + \mathbf{1}\mathbf{1}'\alpha)^{-1} - \mathbf{A}_{22}^{-1} \end{bmatrix}$$

# How to find the value for $\alpha$ ?

$\mu$  is the average BV of genotyped individuals

$$\mu = \frac{1}{n} \mathbf{1}' \mathbf{u}_2$$

Assume traditional BLUP is unbiased.

$\mu$  from either pedigree

$$\mu_p \sim N\left(\mathbf{0}, \frac{1}{n^2} \mathbf{1}' \mathbf{A}_{22} \mathbf{1} \sigma_u^2\right)$$

or single-step

$$\mu_s \sim N\left(\mathbf{0}, \frac{1}{n^2} \mathbf{1}' [\mathbf{G} + \mathbf{1}\mathbf{1}'\alpha] \mathbf{1}\right) \sigma_u^2$$

# How to find the value for $\alpha$ ?

As the  $\mathbf{1}'\mathbf{1}$  are simply summations

from either pedigree

$$\text{Var}(\mu_p) = \sigma_u^2 \frac{1}{n^2} \sum_i \sum_j \mathbf{A}_{22}(i,j)$$

or single-step

$$\text{Var}(\mu_s) = \sigma_u^2 \left( \alpha + \frac{1}{n^2} \sum_i \sum_j \mathbf{G}_{i,j} \right)$$

If we equate both variances of  $\mu$   $\rightarrow$

$$\alpha = \frac{1}{n^2} \left[ \sum_i \sum_j \mathbf{A}_{22}(i,j) - \sum_i \sum_j \mathbf{G}_{i,j} \right]$$

$\alpha$  is simply the difference between means for  $\mathbf{A}_{22}$  and  $\mathbf{G}$

# What does $\alpha$ mean?

$$\alpha = \frac{1}{n^2} \left[ \sum_i \sum_j \mathbf{A}_{22(i,j)} - \sum_i \sum_j \mathbf{G}_{i,j} \right]$$

$\alpha$  accounts for the fact that  $\mathbf{u}_2$  are related through pedigree more than  $\mathbf{G}$  is able to reflect

This is because we do not know base allelic frequencies to construct the 'correct'  $\mathbf{G}$

# What does $\alpha$ mean?



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From Wright's  $F_{ST}$ , another interpretation of  $\alpha$  is possible...

The  $F_{ST}$  can be defined as the mean relationship between gametes in a recent population with respect to an older base population

$$F_{old} = F_{new} + (1 - F_{new}) F_{ST}$$

Powell et al., 2010

# What does $\alpha$ mean ?

$\mathbf{A}_{22}$  involves relationships of genotyped individuals with reference to the base population, and

$\mathbf{G}$  corresponds to relationships within the current population.

Consequently,  $\alpha$  is equal to twice  $F_{ST}$

$$F_{ST} = \frac{1}{2} \text{mean}(\mathbf{A}_{22} - \mathbf{G}) \longrightarrow \mathbf{G}^* = \left(1 - \frac{1}{2}\alpha\right) \mathbf{G} + \mathbf{11}' \alpha$$

# Which G must we use ?

$$\text{AvgD}(\mathbf{G}) = \text{AvgD}(\mathbf{A}_{22})$$

$$\mathbf{G}^* = \frac{\text{trace}(\mathbf{A}_{22})}{\text{trace}(\mathbf{G})} \mathbf{G}$$

2° moment  
(variance) of  $\mathbf{u}_2$

$$\text{sum}(\mathbf{G}) = \text{sum}(\mathbf{A}_{22})$$

$$\mathbf{G}^* = \mathbf{G} + \mathbf{1}\mathbf{1}'\alpha$$

1° moment (mean) of  $\mathbf{u}_2$

Both,

$$\begin{aligned} \text{AvgD}(\mathbf{G}) &= \text{AvgD}(\mathbf{A}_{22}) \\ \text{sum}(\mathbf{G}) &= \text{sum}(\mathbf{A}_{22}) \end{aligned}$$

$$\mathbf{G}^* = \left(1 - \frac{1}{2}\alpha\right) \mathbf{G} + \mathbf{1}\mathbf{1}'\alpha$$

Mean & variance of  $\mathbf{u}_2$   
(assumption of random mating)

# Which G must we use ?

Both,

$$\text{AvgD}(\mathbf{G}) = \text{AvgD}(\mathbf{A}_{22})$$

$$\text{sum}(\mathbf{G}) = \text{sum}(\mathbf{A}_{22})$$

$$\mathbf{G}^* = \left(1 - \frac{1}{2}\alpha\right) \mathbf{G} + \mathbf{1}\mathbf{1}'\alpha$$

Mean & variance of  $\mathbf{u}_2$   
(assumption of random mating)

$$\mathbf{G}^* = a\mathbf{G} + \mathbf{1}\mathbf{1}'b$$

Mean & variance of  $\mathbf{u}_2$  (no random mating)

**preGSf90**

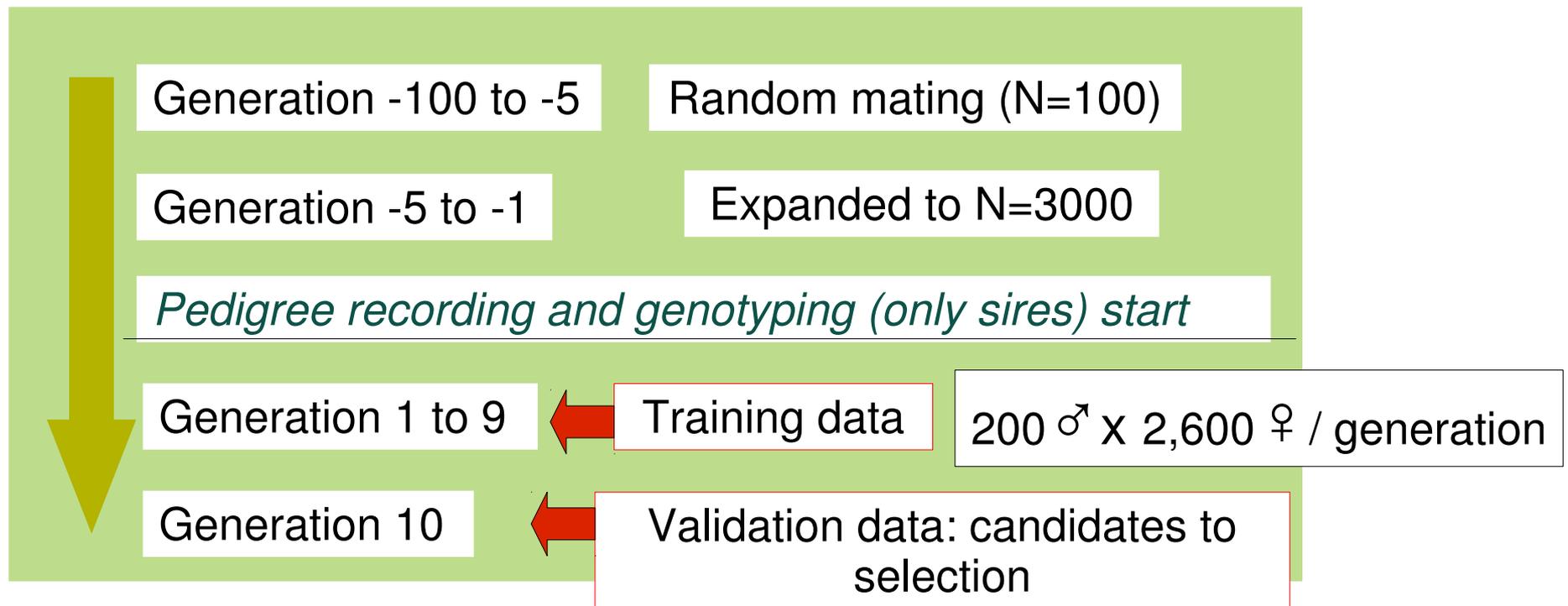
Christensen et al., 2012,

$a = 0.859$  (Christensen et al., 2012)

$a = 0.851$  (Vitezica et al., 2011)

# Simulation (QMSim, 2009)

- ✓ 10 K SNPs - 10 chrom. (10 SNPs/cM)
- ✓ Sex-limited trait ( $h^2$  : 0.5 – 0.3 – 0.05)
- ✓ Selection: phenotype / EBV



# Genetic evaluation

EBVs of the candidates to selection (last generation) were estimated

1) **BLUP<sub>PED</sub>** pedigree-based evaluation

2) **BLUP<sub>DYD</sub>** 2-step procedure

BLUP<sub>PED</sub> -> DYD's for sires -> GBLUP

3) **BLUP<sub>1STEP</sub>** single-step (SS) procedure, standard **G**

4) **BLUP<sub>ALPHA</sub>** SS procedure,  $\mathbf{G}^\dagger = \mathbf{G} + \mathbf{1}\mathbf{1}'\alpha$

5) **BLUP<sub>F<sub>ST</sub></sub>** SS with  $\mathbf{G}^*$  proposed by Powell *et al.*, 2010

# Results

## Means of EBV ( $h^2=0.30$ ) for selection candidates

Prediction method	$P_y$	PEBV
	<b>TBV=0.53 (0.03)</b>	<b>TBV=2.01 (0.15)</b>
<b>BLUP<sub>PED</sub></b>	<b>0.54</b>	<b>2.05</b>
<b>BLUP<sub>DYD</sub></b>	<b>0.17</b>	<b>0.61</b>
<b>BLUP<sub>1STEP</sub></b>	<b>0.29</b>	<b>1.41</b>
<b>BLUP<sub>ALPHA</sub></b>	<b>0.52</b>	<b>2.10</b>
<b>BLUP<sub>Fst</sub></b>	<b>0.52</b>	<b>2.10</b>

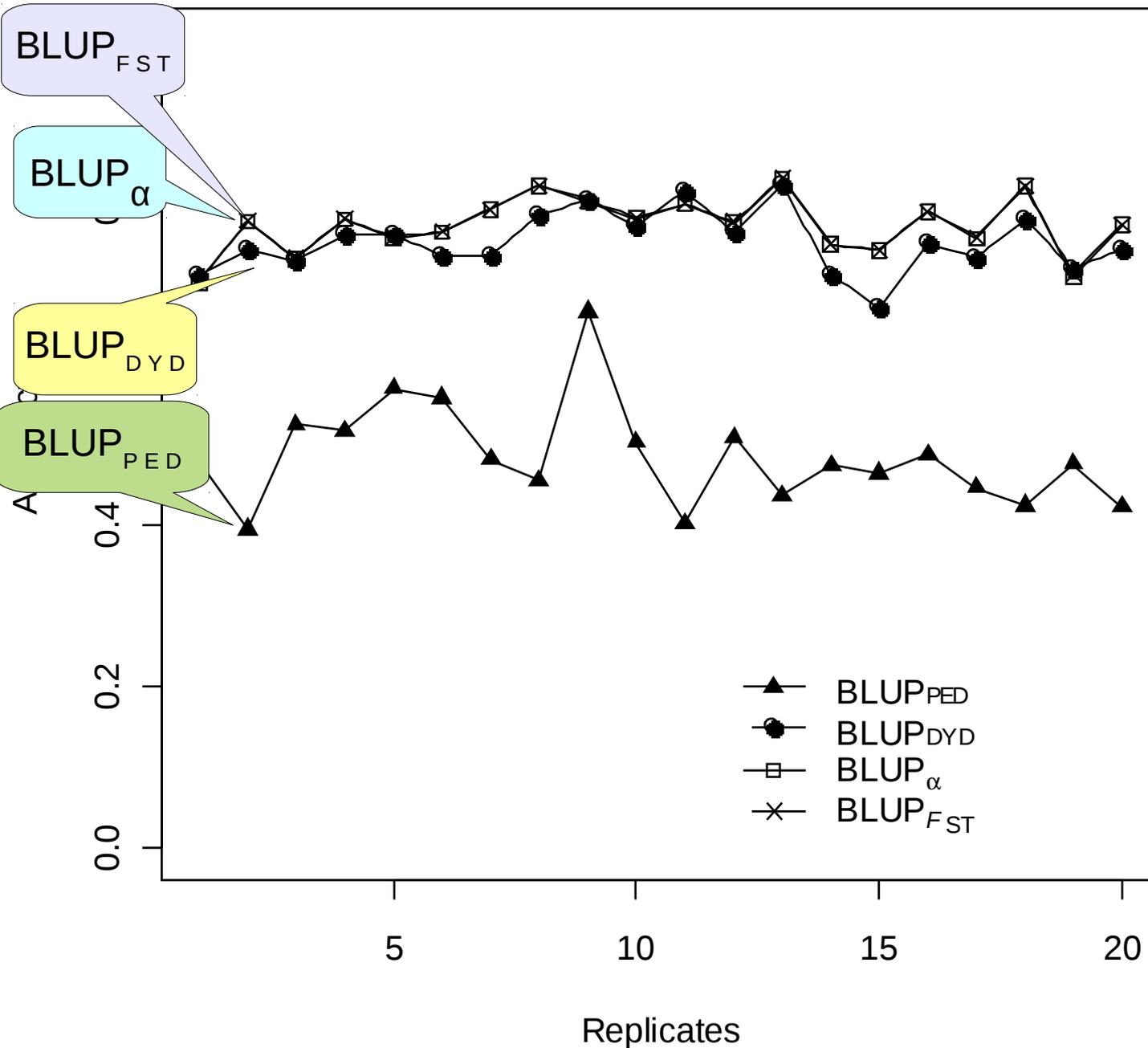
Sd err ~0.04

Sd err ~0.15

# Results



Accuracy of  
different  
genomic  
prediction  
methods across  
20 replicates  
 $h^2=0.30$   
 $P_{EBV}$  selection



# Conclusions



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- ✓ Single-step method with correction (either  $BLUP_{FST}$  or  $BLUP_{Chistensen et al. 2012}$ ) is the recommended method for accounting for bias in genomic predictions
- ✓ Bias is an important concern with 2-step procedure ( $BLUP_{DYD}$ )

What would happen with GEBV in a joint evaluation of multiple lines or breeds ?

# Evaluation of a multi-line broiler chicken population using ssGBLUP

Simeone et al., 2011. JABG.

**Table 1** Descriptions of the phenotypic data for body weight for all animals and genotyped animals in lines A and B, and the multi-line population<sup>1</sup>

Line	No. of records	Mean (SD)
A		
All animals	183 695	24.50 (3.22)
Genotyped animals <sup>2</sup>	3195	25.12 (2.97)
B		
All animals	164 149	23.53 (3.17)
Genotyped animals	3001	23.39 (2.63)
Multi-Line <sup>3</sup>		
All animals	347 844	24.04 (3.24)
Genotyped animals	6196	24.28 (2.94)

Lines were combined to create a multi-line pop



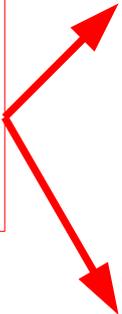
The purpose was to examine changes of GEBVs in analysis of two lines when **G** was constructed with different allele frequencies

**Table 2** Statistics for GEBVs and EBVs estimated for the multi-line population<sup>1</sup>

Population	Mean (G)EBV (SD)	Minimum	Maximum
Line A, traditional	-0.14 (0.47)	-1.32	1.01
Line A, SSP <sup>2</sup> (A)	0.07 (0.59)	-1.71	1.70
Multi (A)	0.07 (0.60)	-1.75	1.76
Multi (B)	-0.12 (0.60)	-1.96	1.54
Multi (AB) <sup>3</sup>	0.13 (0.58)	-1.64	1.70
Multi (0.5)	-0.18 (0.53)	-1.81	1.29
Line B, traditional	-0.28 (0.34)	-1.31	0.69
Line B, SSP (B)	0.00 (0.47)	-1.51	1.31
Multi (A)	-0.25 (0.47)	-1.74	0.93
Multi (B)	0.02 (0.49)	-1.54	1.25
Multi (AB)	-0.06 (0.47)	-1.49	1.11
Multi (0.5)	-0.34 (0.42)	-1.64	0.76

\* Frequencies used to scale G are in parentheses

GEBV predicted using the correct allele frequencies for each line



The means of GEBV changed with different **G**, but animals were ranked appropriately within each line



In ssGBLUP,  $\mathbf{G}$  may be used to evaluate simultaneously multiple lines or breeds, but should be adjusted to obtain proper GEBVs

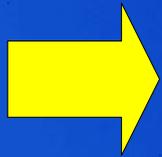
# Multi-breed evaluation with ssGBLUP in New Zealand

Harris et al., 2012.

## Base Adjustment Method

### Computing $G$

- Extended the correction suggested by Vitezica et al. (2011) for multi-breed genomic relationship matrices
- Similar to multi-breed mean adjustments in Harris and Johnson (2010)
- Computationally easier than current adjustments



$$\alpha_{ij} = \mu + \alpha^H p_i^H p_j^H + \alpha^J p_i^J p_j^J + \alpha^X (p_i^H p_j^J + p_j^H p_i^J)$$

Interbull bulletin, n°46, Cork Ireland, May 28 – 31, 2012.